



ORIGINAL RESEARCH

Intra- and inter-rater reliability of Fugl-Meyer Assessment of Lower Extremity early after stroke



Edgar D. Hernández^a, Sandra M. Forero^b, Claudia P. Galeano^b,
Nubia E. Barbosa^b, Katharina S. Sunnerhagen^c, Margit Alt Murphy^{c,*}

^a Departamento del Movimiento Corporal Humano, Universidad Nacional de Colombia, Bogota, Colombia

^b Central Military Hospital of Colombia, Bogota, Colombia

^c Institute of Neuroscience and Physiology, Clinical Neuroscience, Rehabilitation Medicine, Sahlgrenska Academy, University of Gothenburg, Gothenburg, Sweden

Received 3 July 2020; received in revised form 23 November 2020; accepted 2 December 2020

Available online 17 December 2020

KEYWORDS

Item-level reliability;
Leg motor activity;
Scale;
Stroke rehabilitation;
Svensson's method

Abstract

Background: The Fugl-Meyer Assessment of Lower Extremity (FMA-LE) is a widely used and recommended scale for evaluation of post-stroke motor impairment. However, the reliability of the scale has only been established by using parametric statistical methods, which ignores the ordinal properties of the scale.

Objective: To determined intra- and inter-rater reliability of the FMA-LE at item and summed score level early after stroke.

Methods: Sixty patients (mean age 65.9 years, median FMA-LE 29 points) admitted to the hospital due to stroke were included. The FMA-LE was simultaneously, but independently, scored by three experienced and trained physical therapists randomly assigned into pairs, on two consecutive days, between 4 to 9 days post stroke. A rank-based statistical method for paired ordinal data was used to assess the level of agreement and systematic and random disagreements.

Results: The item-level reliability was high (percentage of agreement [PA] $\geq 75\%$). Two items (ankle dorsiflexion during flexor synergy and normal reflex activity) showed some systematic disagreement in intrarater analysis. A satisfactory intrarater reliability (PA $\geq 70\%$) was reached for all summed scores when a 1- or 2-point difference was accepted between ratings.

Conclusion: The FMA-LE is a reliable tool for assessment of motor impairment both within and between raters early after stroke. The scale can be recommended not only for use in Spanish speaking countries, but also internationally. A unified international use of FMA-LE would allow comparison of stroke recovery outcomes worldwide and thereby potentially improve the quality of stroke rehabilitation.

© 2020 The Author(s). Published by Elsevier Editora Ltda. on behalf of Associação Brasileira de Pesquisa e Pós-Graduação em Fisioterapia. This is an open access article under the CC BY license (<http://creativecommons.org/licenses/by/4.0/>).

* Corresponding author at: Per Dubbsgatan 14, Plan 3, 413 45 Gothenburg, Sweden.

E-mail: margit.alt-murphy@neuro.gu.se (M. Alt Murphy).

Introduction

Stroke is the leading cause of disability worldwide.¹ The incidence and global burden of stroke is increasing, particularly in low- and middle-income countries.² Motor impairment is the most prominent impairment after stroke as it affects planning, production, and execution of movements in the contralateral arm and leg.³ Muscle weakness, altered muscle coupling, and co-activation are common motor deficits that often can be more prominent in distal parts of the body particularly when corticospinal descending and ascending neural pathways are involved in the injury.⁴ Motor impairments in lower extremity influence walking ability, walking speed, as well as static and dynamic postural control.⁵⁻¹⁰ The Copenhagen unselected cohort study in stroke showed that approximately 65% of individuals had leg paresis at admission and that 55% had remaining paresis at discharge from an inpatient rehabilitation unit.¹¹ Similarly, about 63% and 36% were either not able to walk or needed assistance at admission and discharge, respectively.¹¹ These numbers indicate that assessment of lower extremity paresis is essential in all stages of stroke to fully understand the mechanisms of motor control and its consequences on walking ability, postural control, and activities of daily living.

The Fugl-Meyer Assessment of Lower Extremity (FMA-LE) is a widely used scale for assessment of motor function after stroke.^{12,13} The scale is recognized as a gold standard and is recommended both for clinical use and research worldwide.^{14,15} The scale includes assessment of reflex activity, voluntary movements within and outside of synergies, ability to perform isolated movement, and coordination. The FMA-LE measures a unidimensional underlying construct, motor impairment, and poses hierarchical properties.^{16,17} This means that the scale is valid for determining level of motor function in people with stroke.

Reliability, cross-sectional and predictive validity, as well as responsiveness of the FMA-LE have been demonstrated by several previous studies.^{15,18} Excellent intra- and inter-rater reliability of the FMA-LE in the sub-acute phase (intraclass correlation coefficient [ICC], 0.95–0.99) and chronic phase (ICC 0.88–0.95) has been reported.¹⁸⁻²⁰ The scoring of each item of the FMA-LE is done at the ordinal level (0–2) and the total score is calculated as a sum-score. The summing of ordinal scores does not result in a number that is valid for making quantitative analysis on reliability or any other comparisons.²¹ Therefore, the results from studies using parametric statistics on ordinal scales, such as ICC, should be interpreted with caution. Furthermore, the reliability of the FMA-LE needs to be established by using methods suited for ordinal data to verify the agreement and not only association between different raters. For predictive purposes the use of single items or sub-scores of longer scales has become of great interest among researchers and clinicians.²²⁻²⁶ Thus, there is a need to establish the intra- and inter-rater reliability at all levels.

The FMA-LE was recently translated into Colombian Spanish following the protocol and manual according to the original English/Swedish version.²⁷ Because the psychometric properties of a scale are dependent on the language, population, and setting, there is a need to assess reliability and validity of the Spanish version of the FMA-LE. Thus, the

aim of this study was to evaluate the intra- and inter-rater reliability of the FMA-LE at the item and summed score level in people early after stroke.

Methods

Participants

This study, investigating intra- and inter-rater reliability, involved a sample of 60 patients consecutively admitted to the Central Military Hospital of Colombia in Bogota during a 17-month period due to stroke. Inclusion criteria were: first event stroke, upper or lower extremity hemiparesis, admitted to the hospital between 4 to 9 days post stroke, age between 18 and 90 years. Exclusion criteria were: other disorders such as blindness, deafness, amputation of lower or upper limb, cerebellar stroke, not able to cooperate in FMA testing due to impaired cognition or other severe medical condition. The severity of the stroke at hospital admission was assessed by the National Institutes of Health Stroke Scale (NIHSS)²⁸ and the disability level at discharge by the Modified Rankin Scale.²⁹

The study protocol was endorsed by the Research Ethics Committee of the Central Military Hospital, Bogota, Colombia (Act No. 9, 12 June 2013) and a signed informed consent was obtained from all participants or their family member. The data collection was conducted between November 2014 and April 2016. The STROBE (Strengthening the Reporting of Observational studies in Epidemiology) guidelines³⁰ and the checklist for reliability evaluation from the consensus-based standards for selection of health status measurement instruments (COSMIN) were followed to ensure the methodological quality of the study.³¹ The statistical rank invariant method used in the current study to determine reliability is not listed in the COSMIN, but it is a valid alternative for determination of reliability in ordinal paired data.³²⁻³⁴ The sample size estimation was based on previous studies using the same statistical methodology.^{35,36}

Fugl-Meyer Assessment of Lower Extremity

The FMA-LE assesses lower extremity motor function including reflex activity, movement within and outside synergy patterns, and speed/coordination.¹³ It comprises 17 items in two subscales: Lower Extremity (E) and Speed/Coordination (F), which are scored on a 3-level ordinal scale (0 points: none; 1 point: partial; 2 points: full). The item scores are then summed. The maximum score for the Lower Extremity Subscale is 28 points and for Speed/Coordination 6 points. The total summed score of 34 points indicates normal function. The protocol used for FMA-LE assessment is available at www.neurophys.gu.se/rehabmed.

Three physical therapists were randomly assigned into pairs of two to perform the assessments. All raters had more than 20 years of clinical experience and underwent training on the FMA-LE prior to the start of the study. All raters were involved in the translation process of the FMA from English into Spanish, which also included joint practical training with guidance of experts and data collection for a previous pilot study.²⁷ The patient's performance on the FMA-LE was simultaneously, but independently, scored

by one pair of raters on two consecutive days. The first assessment was performed between 4 to 9 days post stroke. During the first assessment one of the raters was acting as test leader (i.e. instructing the patient and scoring) and the other as observer (scoring by observing). These roles were switched on the second assessment day. The examiners did not communicate during the testing session or afterwards regarding the scoring. The scoring protocols were stored in sealed envelopes until the data collection was completed.

Statistical analysis

Descriptive statistics were calculated for the background data. Floor and ceiling effects for the FMA-LE were defined as more than 15% of patients receiving the lowest or highest score on the scale.¹⁸

For the intra- and inter-rater reliability, a rank invariant method specially designed for analysis of systematic and non-systematic disagreements in paired ordinal data was used^{32–34} (the software is available at <http://avdic.se/svenssonsmetod.html>). This method was preferred over the weighted kappa, because the latter fails to identify the systematic disagreements and ignores the rank invariant properties of ordinal data.^{37,38} The weighted kappa also assumes that the raters have equal skill level, which means that systematic disagreements are ignored.^{37,38} In addition, the weighted kappa value depends on the choice of weights and is sensitive to the number of categories, which means that the value increases when the number of categories decreases.³⁸

The degree of agreement was determined by using the percentage of agreement (PA) in which agreement $\geq 70\%$ was considered satisfactory.³⁹ For the summed scores, a minimum disagreement in points to reach at least 70% PA was also calculated. The systematic disagreement between raters was expressed as relative position (RP), the relative concentration (RC), and the relative rank variation (RV).³² The RP indicates the extent to which the distribution of scores from an assessment is systematically shifted towards higher or lower categories. The RC shows whether the scores are more or less concentrated towards the central categories of the scale compared to the other assessment. The RP and RC values can vary from -1 to 1, where 0 means no difference between raters. Values within -0.1 and 0.1 were considered negligibly small with reference to clinical relevance, while values outside this range were considered as clinically relevant disagreements.³⁸ The RV indicates disagreement caused by individual variability and varies between 0 and 1 and a value <0.1 means that the difference is negligible. Statistically significant disagreement of RP, RC, and RV was indicated with a 95% confidence interval (95% CI) that did not include the value zero. The statistical software also produced Receiver Operating Characteristic (ROC) curve for each comparison which were used to visually evaluate the detected systematic disagreements. Concave or convex curves indicated disagreement in position and S-shaped curve that raters concentrated their assessment differently on the scale categories. The reliability was considered to be excellent when all systematic and non-systematic disagreements were statistically non-significant within the limits stated above.

Table 1 Demographic and clinical characteristics (n=60).

Characteristics	Value
Age, years, mean \pm SD	65.9 \pm 17.3
Sex, male/female, n (%)	31/29 (52%/48%)
Ischemic/hemorrhagic stroke, n (%)	55/5 (93%/7%)
Right/left hemiparesis, n (%)	33/27 (55%/45%)
Thrombolysis, n	8
Hospitalization days, mean \pm SD	12 \pm 10
Days post stroke to first assessment, mean \pm SD	5.95 \pm 2.73
Modified Rankin Scale, median (Q1–Q3)	2 (1–4)
0–2 Mild or non-significant disability, n	35
3–5 Moderate to severe disability, n	25
NIHSS Scale, median (Q1–Q3)	5 (3–10)
Mild 0–5, n	25
Moderate to severe 6–24, n	22
Patients without NIHSS scorings, n	13
Discharged from hospital	
Home, n	56
Homecare, n	1
Intermediate care, n	1
Died in hospital, n	2
Fugl Meyer Assessment of Lower Extremity (FMA-LE)	
FMA-LE, 1st occasion, median (Q1–Q3)	29 (26–31)
FMA-LE, 2nd occasion, median (Q1–Q3)	29.5 (27–31)

Abbreviations: NIHSS, National Institutes of Health Stroke Scale.

Results

Out of 105 eligible patients, 45 were excluded due to: limited ability to follow test instructions (n=21), cerebellar stroke (n=8), severe multi-impairment (n=6), discharged (n=5), prior stroke (n=4), deceased (n=1). All 60 patients (31 men and 29 women, mean age of 65.9 years) included in the study were able to perform the FMA-LE (Table 1). The majority (93%) had ischemic stroke and 7% had haemorrhagic stroke. The FMA-LE scores of the study group ranged from 4 to 34 points. The FMA-LE showed no floor or ceiling effect (9 patients received full score of 34 points).

Intrarater reliability

The intrarater reliability was calculated separately for all three raters. At the item level the PA across all raters was above 75% for all tested items (Table 2). The rank invariant analysis of agreement revealed statistically significant disagreement of RP (≥ 0.1) for the ankle dorsiflexion within synergies in supine position (E.II) and for test of normal reflex activity (E.V) in one of the raters (Table 3). The ankle dorsiflexion in supine (E.II) and in standing position (E.IV) showed also a tendency towards non-negligible disagreements (asymmetric 95% CI) in concentration and position, respectively. All these disagreements were positive, which

Table 2 Percentage of agreement (PA%) within each rater (A, B, and C) and between test occasions.

E. Lower extremity	Intrarater agreement (PA %)			Interrater agreement (PA %)	
	Rater A n = 40	Rater B n = 38	Rater C n = 38	Test occasion 1 n = 60	Test occasion 2 n = 60
I. Reflex activity					
Flexors	97	97	97	100	100
Extensors	100	100	97	100	100
II. Within synergies, supine					
Hip flexion	95	94	92	95	97
Knee flexion	97	100	97	100	100
Ankle dorsiflexion	95	88	89 ^a	92	95
Hip extension	92	97	97	100	98
Hip adduction	97	100	94	98	98
Knee extension	97	100	100	97	97
Ankle plantar flexion	95	88	89	98	93
<i>SUM E II, range 0–14 points</i>	90	79	83	88	90
III. Mixed synergies, sitting					
Knee flexion	95	97	92	98	95
Ankle dorsiflexion	90	91	86	93	95
<i>SUM E III, range 0–4 points</i>	90	91	83	92	93
IV. Little or no synergies, standing					
Knee flexion to 90°	92	82	81	95	93
Ankle dorsiflexion	77	76	86	88	95
<i>SUM E IV, range 0–4 points</i>	69 ^a	68	72 ^a	85	92
<i>SUM E IV, 1-point difference</i>	100	95	94	–	–
V. Normal reflex activity					
Knee flexors, patellar, Achilles	92	91	83 ^a	100	100
<i>SUM E, range 0–28 points</i>	62	62	53 ^a	78	85
<i>SUM E, 1-point difference accepted</i>	87	79	69	–	–
<i>SUM E, 2-point difference accepted</i>	94	86	82	–	–
F. Coordination/speed					
Tremor	87	100	97	93	92
Dysmetria	85	91	83	90	92
Time	77	76	75	95	97
<i>SUM F, range 0–6 points</i>	62	68	64	83	87
<i>SUM F, 1-point difference accepted</i>	90	97	91	–	–
TOTAL E-F, range 0–34 points	51	47	50 ^a	75	80
TOTAL E-F, 1-point difference accepted	77	76	69	–	–
TOTAL E-F, 2-point difference accepted	92	82	80	–	–

Abbreviations: PA, percentage of agreement; RP, relative position.

^a Statistically significant disagreement where the absolute value of RP is ≥ 0.1 and the 95% confidence interval does not include 0 are marked in bold.

indicates that a higher category was systematically more frequently used at the second occasion for these items sub scores or total scores. No individual disagreement measured as random variance was noted across raters.

At the summed score level (Table 2), 79%–100% agreement was reached for movements performed within and mixed synergies (E.II and E.III), and 62%–72% was reached for movements performed with little or no synergy (E.IV) and coordination/speed (F). A disagreement in relative position was revealed for the sum-score of little or no synergy (E.IV) (Table 3). For the summed score E including all motor items (possible maximum

score of 28 points), the agreement within raters varied between 53% and 62% (Table 2). When all items were summed to a total score E-F (maximum score of 34 points), the agreement varied between 47% and 51%. The lower PA values in the summed scores were expected because the number of possible categories is larger. However, 69%–87% PA was reached for sum-score E, and for total sum-score E-F when a 1-point difference between test occasions was accepted. Thus, a satisfactory intrarater reliability at sum-score levels was reached when 1- or 2-points difference between test-occasions was accepted.

Table 3 The rank invariant analysis of intrarater agreement within raters A, B, and C.

E. Lower extremity	Rater A		Rater B		Rater C	
	RP (95% CI)	RC (95% CI)	RP (95% CI)	RC (95% CI)	RP (95% CI)	RC (95% CI)
I. Reflex activity						
Flexors	0.03 (-0.02, 0.07)	-	0.03 (-0.03, 0.09)	-	0.03 (-0.03, 0.08)	-
Extensors	0	-	0	-	0.03 (-0.03, 0.08)	-
SUM E I (0-4 points)	0.03 (-0.02, 0.07)	-	0.03 (-0.03, 0.09)	-	0.03 (-0.03, 0.08)	-
II. Within synergies, supine position						
Hip flexion	0.05 (-0.02, 0.11)	-0.04 (-0.09, 0.02)	0 (-0.08, 0.08)	0 (-0.03, 0.03)	0.03 (-0.06, 0.11)	-0.01 (-0.07, 0.04)
Knee flexion	0.02 (-0.02, 0.07)	-0.02 (-0.06, 0.02)	0	0	0.03 (-0.02, 0.08)	-0.02 (-0.07, 0.02)
Ankle dorsal flexion	0.04 (-0.02, 0.10)	-0.03 (-0.08, 0.02)	0.04 (-0.04, 0.11)	-0.10 (-0.21, 0.01) ^b	0.10 (0.01, 0.12) ^a	-0.06 (-0.14, 0.02)
Hip extension	0 (-0.07, 0.06)	-0.03 (-0.12, 0.06)	0 (-0.01, 0)	-0.03 (-0.09, 0.03)	0.03 (-0.02, 0.08)	-0.02 (-0.07, 0.02)
Hip adduction	0.02 (-0.02, 0.07)	-0.03 (-0.08, 0.03)	0	0	0.05 (-0.02, 0.12)	-0.04 (-0.11, 0.03)
Knee extension	0 (-0.01, 0)	-0.03 (-0.08, 0.03)	0	0	0	0
Ankle plantar flexion	0.04 (-0.02, 0.10)	-0.03 (-0.08, 0.02)	0.02 (-0.07, 0.11)	-0.05 (-0.14, 0.04)	0.05 (-0.05, 0.15)	-0.03 (-0.09, 0.03)
SUM E II (0-14 points)	0.04 (-0.01, 0.10)	0	0.06 (-0.05, 0.18)	0	0.09 (0, 0.19)	0
III. Mixed synergies, sitting position						
Knee flexion	0 (-0.06, 0.06)	0 (-0.05, 0.05)	-0.03 (-0.09, 0.03)	-0.02 (-0.07, 0.03)	0 (-0.07, 0.07)	0 (-0.08, 0.08)

Table 3 (Continued)

E. Lower extremity	Rater A		Rater B		Rater C	
	RP (95% CI)	RC (95% CI)	RP (95% CI)	RC (95% CI)	RP (95% CI)	RC (95% CI)
Ankle dorsiflexion	0 (-0.09, 0.09)	0 (-0.06, 0.06)	-0.01 (-0.08, 0.07)	-0.03 (-0.11, 0.04)	-0.01 (-0.10, 0.09)	-0.03 (-0.12, 0.06)
SUM E III (0–4 points)	0 (-0.09, 0.09)	0 (-0.06, 0.06)	-0.01 (-0.08, 0.06)	-0.05 (-0.16, 0.07)	0 (-0.10, 0.10)	0 (-0.10, 0.09)
IV. Little or no synergy, standing position						
Knee flexion to 90°	0.06 (-0.01, 0.13)	-0.06 (-0.13, 0.01)	0.05 (-0.07, 0.17)	-0.04 (-0.14, 0.07)	0.07 (-0.04, 0.19)	-0.05 (-0.16, 0.07)
Ankle dorsiflexion	0.11 (-0.01, 0.23) ^b	-0.09 (-0.21, 0.03)	-0.05 (-0.19, 0.10)	-0.06 (-0.17, 0.05)	0.07 (-0.01, 0.16)	-0.05 (-0.17, 0.07)
SUM E IV (range 0–4 points)	0.12 (0.02, 0.23) ^a	-0.05 (-0.16, 0.06)	0.01 (-0.11, 0.13)	-0.06 (-0.19, 0.07)	0.11 (0.01, 0.21) ^a	-0.05 (-0.21, 0.11)
V. Normal reflex activity						
Knee flexors, patellar, Achilles	0.03 (-0.06, 0.11)	0 (-0.07, 0.13)	0.03 (-0.07, 0.13)	0 (-0.05, 0.19)	0.17 (0.05, 0.29) ^a	0 (0.03, 0.23) ^a
SUM E (0–28 points)	0.06 (0, 0.12)	0 (-0.05, 0.19)	0.07 (-0.05, 0.19)	0 (-0.03, 0.25) ^b	0.13 (-0.10, 0.11)	0 (-0.13, 0.14)
F. Coordination/speed						
Tremor	-0.09 (-0.18, 0)	0.09 (-0.02, 0.19)	0 (-0.01, 0.15)	0 (-0.06, 0.10)	-0.03 (-0.08, 0.02)	0.02 (-0.02, 0.05)
Dysmetria	0.04 (-0.06, 0.13)	0.01 (-0.08, 0.11)	0.07 (-0.01, 0.15)	0.02 (-0.06, 0.10)	0.09 (-0.02, 0.20)	0.01 (-0.08, 0.09)
Time	0 (-0.11, 0.10)	0.05 (-0.10, 0.21)	0.09 (-0.04, 0.22)	0.07 (-0.07, 0.21)	-0.06 (-0.20, 0.09)	0 (-0.12, 0.13)
SUM F (0–6 points)	-0.03 (-0.10, 0.05)	0.04 (-0.06, 0.15)	0.09 (-0.01, 0.19)	0.11 (-0.03, 0.25) ^b	0.01 (-0.10, 0.11)	0 (-0.13, 0.14)
Total E-F (0–34 points)	0.05 (-0.01, 0.11)	0 (-0.02, 0.22) ^b	0.10 (-0.02, 0.22) ^b	0 (-0.02, 0.22) ^b	0.11 (0.02, 0.21) ^a	0 (0.02, 0.21) ^a

Abbreviations: PA, percentage of agreement, RP, relative position; RC, relative concentration; CI, confidence interval.

Absolute values of RP/RC ≤ 0.01 are assigned value 0.

^a Statistically significant disagreement (absolute value of RP/RC ≥ 0.1 and 95%CI does not include 0, marked in bold).

^b Tendency towards a non-negligible disagreement (absolute value of RP/RC ≥ 0.1 and asymmetric 95%CI around 0).

Interrater reliability

The FMA-LE scores for each item showed high level of agreement (all above 88%) between raters at both test occasions (**Table 2**). The PA for summed scores of each section varied between 83% and 100%. PA for the summed score E was between 78% and 85%, and for the total sum-score E-F between 75% and 80%. Disagreements were negligible or not statistically significant (**Table 4**). No individual disagreements measured as random variance was noted across raters.

Discussion

This study demonstrated that the FMA-LE is a reliable clinical instrument for evaluation of motor function after stroke. Apart from two items in which systematic disagreements were observed, the item level intra- and inter-rater reliability was excellent. The interrater reliability at the summed score levels was excellent, although, in the intrarater analysis, a shift towards higher scores at the second test occasion was observed in few cases. The level of agreement was satisfactory for the summed sub-scores and the total score when 1- or 2-point difference between ratings was accepted.

The results of this study confirm the excellent intra- and inter-rater reliability, at item and summed score level, previously shown for the Fugl-Meyer Assessment of Upper Extremity (FMA-UE).³⁶ Furthermore, the item level reliability of FMA-LE was even higher compared to upper extremity assessment. In FMA-LE, only two items, the ankle dorsiflexion during flexor synergy and normal reflex activity, demonstrated statistically significant systematic disagreement within-raters when assessed one day apart early after stroke. Similar to the reliability study of the FMA-UE,³⁶ a systematic shift towards higher scores at the second test occasion was observed, which is indicative for possible spontaneous recovery at this early stage of stroke. To improve reliability of items and the sum scores that showed systematic disagreements, clearer guideline and training might be needed.

The intra- and inter-rater reliability of the FMA-LE has shown to be excellent using parametric statistical analysis.^{18–20} In a study with similar sample size to ours, an ICC score of 0.95 was reported for the FMA-LE in a chronic stroke population.¹⁸ Intrarater reliability ICC was as high as 0.99 among expert raters and 0.91 when experts were compared to trained raters.¹⁹ Equally high ICC was reported for four physical therapist who underwent joint training for the FMA-LE.²⁰ The results from our study extend these findings further by showing that when the scale was analysed as an ordinal scale, the interrater reliability at the total score level was excellent (PA 80% and no observed disagreements). The intrarater reliability of the total FMA-LE score demonstrated some systematic shift towards higher scores at the second test occasion. Correspondingly, only 50% of ratings had the exact same score at both occasions. Our results showed, however, that when a 1-point difference of the total score was accepted, a 69% agreement was reached. Likewise, a 2-point difference resulted in an 80% agreement between ratings. These absolute values on the expected variance of the summed scores between

ratings can be useful for clinicians when making distinction between real improvement and random measurement errors in repeated assessments. The absolute disagreement in the FMA-LE scores was clearly under the reported minimal important difference,^{18,19} which confirms the stability of this scale. It is important to note that the results from this study are primarily applicable for the hospitalized patients in the acute and subacute phase of stroke.⁴⁰

The item-level reliability has previously only been determined for the FMA-UE,³⁶ in which an agreement between 79% to 100% was found between and within raters. These results are analogous to the found agreement for the FMA-LE in this study. The item-level reliability, in particular of the motor items of FMA-LE, was also high in a small sample of patients with subacute stroke included in a transcultural validation of FMA scale into Italian language.⁴¹ The item-level reliability is important to establish together with the reliability of the summed scores. The use of single items of the FMA-UE scale has been of great interest in prediction of motor recovery post stroke.^{22,24} Our results demonstrate that most of the single items of FMA-LE can reliably be used in repeated measures. This opens an opportunity to evaluate the potential of single items or combination of a set of items as potential indices for prediction of motor or functional outcome poststroke.

Strengths and limitations

The strength of this study is the large sample size and the consecutive inclusion of a representative cohort of patients admitted to acute hospital care and rehabilitation after a first-time stroke. The initial motor impairment, assessed 4–9 days post stroke onset, showed that most of the patients had moderate to mild lower extremity impairment, with the FMA-LE scores covering the entire range of the scale without showing floor or ceiling effects. The characteristics of the study sample, however, needs to be considered when generalizing the results.

The COSMIN checklist recommends the use of weighted kappa for analysis of reliability in ordinal data. In the current study, a rank invariant method especially designed for paired ordinal data was used. This choice was based on the fact that, different from weighted kappa, the rank invariant method can identify systematic disagreement and considers the invariant properties of ordinal data. Weighted kappa assumes that the raters have equal skill level, which means that systematic disagreements are ignored.^{37,38} Additionally, the weighted kappa value is sensitive to the choice of weights and the number of categories, meaning that a higher value can be achieved when the number of categories is low.³⁸ Thus, the rank-based method used in the current study have some advantages compared to the weighted kappa statistics.

The relatively short time interval, 1 day, between the first and second assessment used for intrarater reliability, might have caused a recall bias for the raters. This was, however, considered to have lesser influence than a possible improvement of motor function at this early stage of stroke. Indeed, the results suggest that even a one day interval might have been too long for establishing intrarater reliability early after stroke, since a positive systematic shift was observed.

Table 4 The rank invariant analysis of interrater agreement between test occasions.

E. lower extremity	Test occasion 1		Test occasion 2	
	RP (95% CI)	RC (95% CI)	RP (95% CI)	RC (95% CI)
I. Reflex activity				
Flexors, extensors	0	0	0	-
II. Within synergies, supine position				
Hip flexion	-0.05 (-0.10, 0)	0.02 (-0.01, 0.06)	0.03 (-0.01, 0.07)	-0.02 (-0.04, 0.01)
Knee flexion	0	0	0	0
Ankle dorsal flexion	-0.04 (-0.10, 0.02)	0 (-0.06, 0.05)	-0.03 (-0.06, 0.02)	0.04 (-0.01, 0.09)
Hip extension	0	0	0.02 (-0.01, 0.04)	-0.02 (-0.05, 0.01)
Hip adduction	0.02 (-0.01, 0.04)	-0.01 (-0.04, 0.01)	0.02 (-0.01, 0.04)	-0.02 (-0.05, 0.01)
Knee extension	-0.01 (-0.04, 0.02)	0.03 (-0.01, 0.07)	0 (-0.04, 0.04)	0 (-0.03, 0.03)
Ankle plantar flexion	-0.02 (-0.04, 0.01)	0.01 (-0.01, 0.03)	-0.01 (-0.06, 0.04)	0.03 (-0.02, 0.08)
SUM E II, 0–14points	-0.06 (-0.11, -0.01) ^a	0	-0.01 (-0.05, 0.04)	0
III Mixed synergies, sitting position				
Knee flexion	-0.02 (-0.05, 0.01)	0.01 (-0.01, 0.03)	-0.02 (-0.07, 0.04)	0.01 (-0.03, 0.05)
Ankle dorsiflexion	-0.05 (-0.10, 0)	0.01 (-0.05, 0.06)	0.03 (-0.02, 0.07)	-0.04 (-0.09, 0.01)
SUM E III, 0–4 points	-0.05 (-0.09, 0)	0.01 (-0.05, 0.07)	0.03 (-0.02, 0.07)	-0.04 (-0.09, 0)
IV. Little or no synergy, standing position				
Knee flexion to 90°	0 (-0.06, 0.07)	-0.03 (-0.05, 0.01)	0.03 (-0.03, 0.08)	0.01 (-0.04, 0.05)
Ankle dorsiflexion	-0.01 (-0.01, 0.06)	-0.02 (-0.08, 0.04)	-0.01 (-0.06, 0.03)	0.01 (-0.03, 0.05)
SUM E IV, 0–4 points	0 (-0.07, 0.07)	-0.06 (-0.12, 0.01)	0 (-0.04, 0.05)	-0.01 (-0.05, 0.03)
V. Normal reflex activity				
Knee flexors, patellar, Achilles	0	0	0	0
SUM E, range 0–28 points	-0.03 (0.08, 0.01) ^a	0	0.02 (-0.02, 0.05)	0
F. coordination/speed				
Tremor	-0.02 (-0.08, 0.04)	-0.01 (-0.06, 0.04)	0.02 (-0.05, 0.09)	0.03 (-0.01, 0.07)
Dysmetria	-0.01 (-0.06, 0.05)	0.03 (-0.04, 0.10)	0 (-0.06, 0.06)	0.05 (-0.01, 0.11)
Time	0.02 (-0.03, 0.07)	0.01 (-0.04, 0.05)	0 (-0.04, 0.04)	0 (-0.04, 0.04)
SUM F, 0–6 points	0 (-0.05, 0.05)	0.03 (-0.04, 0.11)	-0.01 (-0.05, 0.04)	0.02 (-0.05, 0.09)
Total E–F, range 0–34 points	-0.03 (-0.07, 0.01)	0	0 (-0.03, 0.04)	0

Abbreviations: PA, percentage of agreement, RP, relative position, RC, relative concentration.

Absolute values of RP/RC ≤ 0.01 are signed value 0.

^a Statistically significant but negligible disagreement (absolute RP/RC < 0.1 and 95%CI does not include 0).

This shift could also be caused by the learning effect, occurring both in patients and raters when getting familiar with testing procedures at the second occasion. To minimize the bias in scorings, prior training is needed. In this study, all three physical therapists conducting the assessments had undergone joint training prior to data collection. Additionally, each rater had extensive clinical experience with stroke rehabilitation. Training together with clear protocols and

instructions, preferably in the form of an instructional video, are needed to ensure high level reliability.

Conclusions

The FMA-LE showed excellent intra- and inter-rater reliability in a representative cohort of patients early after stroke in the inpatient rehabilitation setting. The FMA-LE can be

recommended as a reliable tool for assessment of motor impairment both at item- as well as summed score levels. A wider use of the FMA-LF both in Spanish speaking countries and worldwide in the inpatient care settings would strengthen the reporting of stroke outcomes and make comparisons between regions and countries possible, thereby improving the quality of care.

Conflicts of interest

The author declares no conflicts of interest.

Acknowledgments

The authors wish to express their appreciation to the Central Military Hospital of Colombia, Universidad Nacional de Colombia and the patients who participated in the study; and the late Nancy Stella Landinez Parra who was one of the initiators of the study, and took active part of the study planning and data collection.

This project was funded by the Central Military Hospital [Research Project No. 2013059]; the Strengthening Established Partnerships 2017 at University of Gothenburg; the Swedish state under the agreement between the Swedish government and the country councils, the ALF-agreement [ALFGBG-775561, ALFGBG-718711]; Swedish Research Council [VR2017-00946]. The funding bodies had no role in any part of this study.

References

1. Katan M, Luft A. Global burden of stroke. *Semin Neurol*. 2018;38:208–211.
2. Group GBDNDC. Global, regional, and national burden of neurological disorders during 1990–2015: a systematic analysis for the Global Burden of Disease Study 2015. *Lancet Neurol*. 2017;16:877–897.
3. Langhorne P, Coupar F, Pollock A. Motor recovery after stroke: a systematic review. *Lancet Neurol*. 2009;8:741–754. Research Support, Non-U.S. Gov't, Review 2009/07/18.
4. Sanchez N, Acosta AM, Lopez-Rosado R, et al. Lower extremity motor impairments in ambulatory chronic hemiparetic stroke: evidence for lower extremity weakness and abnormal muscle and joint torque coupling patterns. *Neurorehabil Neural Repair*. 2017;31:814–826.
5. Geurts AC, de Haart M, van Nes IJ, et al. A review of standing balance recovery from stroke. *Gait Posture*. 2005;22:267–281, 2005/10/11.
6. Higginson JS, Zajac FE, Neptune RR, et al. Muscle contributions to support during gait in an individual with post-stroke hemiparesis. *J Biomech*. 2006;39:1769–1777, 2005/07/28.
7. Wong SS, Yam MS, Ng SS. The Figure-of-Eight walk test: reliability and associations with stroke-specific impairments. *Disabil Rehabil*. 2013;35:1896–1902, 2013/04/23.
8. Burke E, Dobkin BH, Noser EA, et al. Predictors and biomarkers of treatment gains in a clinical stroke trial targeting the lower extremity. *Stroke*. 2014;45:2379–2384, 2014/07/30.
9. Kwan MS, Hassett LM, Ada L, et al. Relationship between lower limb coordination and walking speed after stroke: an observational study. *Braz J Phys Ther*. 2019;23:527–531.
10. Aguiar LT, Camargo LBA, Estarline LD, et al. Strength of the lower limb and trunk muscles is associated with gait speed in individuals with sub-acute stroke: a cross-sectional study. *Braz J Phys Ther*. 2018;22:459–466.
11. Jorgensen HS, Nakayama H, Raaschou HO, et al. Recovery of walking function in stroke patients: The Copenhagen Stroke Study. *Arch Phys Med Rehabil*. 1995;76:27–32.
12. Duncan Millar J, van Wijck F, Pollock A, et al. Outcome measures in post-stroke arm rehabilitation trials: do existing measures capture outcomes that are important to stroke survivors, carers, and clinicians? *Clin Rehabil*. 2019;33: 737–749.
13. Fugl-Meyer AR, Jaasko L, Leyman I, et al. The post-stroke hemiplegic patient. 1. A method for evaluation of physical performance. *Scand J Rehabil Med*. 1975;7:13–31.
14. Kwakkel G, Lannin NA, Borschmann K, et al. Standardized measurement of sensorimotor recovery in stroke trials: consensus-based core recommendations from the Stroke Recovery and Rehabilitation Roundtable. *Int J Stroke*. 2017;12:451–461.
15. Bushnell C, Bettger JP, Cockroft KM, et al. Chronic stroke outcome measures for motor function intervention trials: expert panel recommendations. *Circ Cardiovasc Qual Outcomes*. 2015;8:S163–169.
16. Crow JL, Harmeling-van der Wel BC. Hierarchical properties of the motor function sections of the Fugl-Meyer assessment scale for people after stroke: a retrospective study. *Phys Ther*. 2008;88:1554–1567.
17. Crow JL, Kwakkel G, Bussmann JB, et al. Are the hierarchical properties of the Fugl-Meyer assessment scale the same in acute stroke and chronic stroke? *Phys Ther*. 2014;94: 977–986.
18. Hsueh IP, Hsu MJ, Sheu CF, et al. Psychometric comparisons of 2 versions of the Fugl-Meyer Motor Scale and 2 versions of the Stroke Rehabilitation Assessment of Movement. *Neurorehabil Neural Repair*. 2008;22:737–744, 2008/07/23.
19. Sullivan KJ, Tilson JK, Cen SY, et al. Fugl-Meyer assessment of sensorimotor function after stroke: standardized training procedure for clinical practice and clinical trials. *Stroke*. 2011;42:427–432, 2010/12/18.
20. Duncan PW, Propst M, Nelson SG. Reliability of the Fugl-Meyer assessment of sensorimotor recovery following cerebrovascular accident. *Phys Ther*. 1983;63:1606–1610, 1983/10/01.
21. Laver Fawcett A. *Principles of Assessment and Outcome Measurement for Occupational Therapists and Physiotherapists: Theory, Skills and Application*. Wiley; 2013.
22. Nijland RH, van Wegen EE, Harmeling-van der Wel BC, et al. Presence of finger extension and shoulder abduction within 72 hours after stroke predicts functional recovery: early prediction of functional outcome after stroke: the EPOS cohort study. *Stroke*. 2010;41:745–750.
23. Smith MC, Barber PA, Stinear CM. The TWIST algorithm predicts time to walking independently after stroke. *Neurorehabil Neural Repair*. 2017;31:955–964.
24. Ghaziani E, Coupe C, Siersma V, et al. Easily conducted tests during the first week post-stroke can aid the prediction of arm functioning at 6 months. *Front Neurol*. 2019;10:1371, 2020/01/30.
25. Cioncoloni D, Veerbeek JM, van Wegen EE, et al. Is it possible to accurately predict outcome of a drop-foot in patients admitted to a hospital stroke unit? *Int J Rehabil Res*. 2013;36:346–353, 2013/04/13.
26. Veerbeek JM, Van Wegen EE, Harmeling-Van der Wel BC, et al. Is accurate prediction of gait in nonambulatory stroke patients possible within 72 hours poststroke? The EPOS study. *Neurorehabil Neural Repair*. 2011;25:268–274.
27. Barbosa NE, Forero SM, Galeano CP, et al. Translation and cultural validation of clinical observational scales — the Fugl-Meyer assessment for post stroke sensorimotor function in Colombian Spanish. *Disabil Rehabil*. 2018;1–7.

28. Dancer S, Brown AJ, Yanase LR. National institutes of health stroke scale in plain english is reliable for novice nurse users with minimal training. *J Emerg Nurs.* 2017;43:221–227.
29. van Swieten JC, Koudstaal PJ, Visser MC, et al. Interobserver agreement for the assessment of handicap in stroke patients. *Stroke.* 1988;19:604–607.
30. Vandebroucke JP, von Elm E, Altman DG, et al. Strengthening the Reporting of Observational Studies in Epidemiology (STROBE): explanation and elaboration. *Ann Intern Med.* 2007;147:W163–W194.
31. Mokkink LB, Terwee CB, Knol DL, et al. The COSMIN checklist for evaluating the methodological quality of studies on measurement properties: a clarification of its content. *BMC Med Res Methodol.* 2010;10:22. Research Support, Non-U.S. Gov't 2010/03/20.
32. Svensson E, Schillberg B, Kling AM, et al. Reliability of the balanced inventory for spinal disorders, a questionnaire for evaluation of outcomes in patients with various spinal disorders. *J Spinal Disord Tech.* 2012;25:196–204.
33. Ardic A, Svensson E. *Svenssons Method (Version 1.1).* Örebro: Interactive Software Supporting Svenssons Method; 2010 [Accessed 26 November 2018].
34. Svensson E, Holm S. Separation of systematic and random differences in ordinal rating scales. *Stat Med.* 1994;13:2437–2453.
35. Nordin A, Alt Murphy M, Danielsson A. Intra-rater and inter-rater reliability at the item level of the Action Research Arm Test for patients with stroke. *J Rehabil Med.* 2014;46:738–745.
36. Hernandez ED, Galeano CP, Barbosa NE, et al. Intra- and inter-rater reliability of Fugl-Meyer Assessment of Upper Extremity in stroke. *J Rehabil Med.* 2019;51:652–659, 2019/08/27.
37. Svensson E. Guidelines to statistical evaluation of data from rating scales and questionnaires. *J Rehabil Med.* 2001;33:47–48.
38. Svensson E. Different ranking approaches defining association and agreement measures of paired ordinal data. *Stat Med.* 2012;31:3104–3117.
39. Kazdin AE. Artifact, bias, and complexity of assessment: the ABCs of reliability. *J Appl Behav Anal.* 1977;10:141–150.
40. Bernhardt J, Hayward KS, Kwakkel G, et al. Agreed definitions and a shared vision for new standards in stroke recovery research: the stroke recovery and rehabilitation roundtable taskforce. *Neurorehabil Neural Repair.* 2017;31:793–799, 2017/09/25.
41. Cecchi F, Carrabba C, Bertolucci F, et al. Transcultural translation and validation of Fugl–Meyer assessment to Italian. *Disabil Rehabil.* 2020:1–6.